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# Introduction to the statistical analysis of single cell transcritomic data

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### Outline

- 1 Brief Presentation of single cell sequencing
- 2 Differential Expression Analysis for sequencing data
- **3** Differential Expression Analysis for single cell data
- 4 Linear Dimension reduction and data visualization
- **5** Alternatives to PCA, non linear embedding methods

#### 6 Conclusions

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# Cell biology revolution

- The cell has been discovered in the 17th century
- Cells are the basic unit of structure and function in living organisms
- Physiology emerges as the meta-cellular science (interaction between cells)



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Cell sorting and the investigation of between-cell variations

- Development of monoclonal antibodies (~ 70s)
- Cell sorting by fluoresence (FACS)
- Cell specific proteins, DNA content
- Limited to a few markers



### Cell biology is going molecular

#### Visualization of Single RNA Transcripts in Situ

Andrea M. Femino, Fredric S. Fay,† Kevin Fogarty, Robert H. Singer\*

Fluorescence in situ hybridization (FISH) and digital imaging microscopy were modified to allow detection of single RNA molecules. Oligodeoxynucleotide probes were synthesized with five fluorocromes per molecule, and the light emitted by a single probe was calibrated. Points of light in exhaustively deconvolved images of hybridized cells gave fluorescent intensities and distances between probes consistent with single messenger RNA molecules. Analysis of  $\beta$ -actin transcription sites after serum induction revealed synchronous and cyclical transcription from single genes. The rates of transcription initiation and termination and messenger RNA processing could be determined by positioning probes along the transcription nuit. This approach extends the power of FISH to yield quantitative molecular information on a single cell.

30–50 probes labeled on 3' end Positions selected to optimize GC content

Open reading frame of target mRNA

SCIENCE • VOL. 280 • 24 APRIL 1998 • www.sciencemag.org

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# Cell biology goes genome-wide

- Classify cells into distinct cell types
- Shape, location, interactions, function
- Recent technological breakthroughs allow the molecular characterization of cells



### The single-cell rule

if IT exists, there is a single-cell version of IT (sooner or later)

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### The RT-qPCR version



- Precise for a small number of genes
- not too expensive (100 cells)
- pros and cons of RT-qPCR are well known

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# The Sequencing version [6]



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# The UMI version [6]



- up to 20,000 genes analyzed
- static snapshot
- expensive (for hundreds of cells)

## The split version [12]



- up to 20,000 genes analyzed
- for millions of cells
- cheap

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### A timeline: technologies [15]



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# The Moore's law of single cell [10]



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# The human cell Atlas project

- comprehensive reference catalog of all human cells
- use stable properties, transient features, locations and abundances.
- describe each human cell by a defined set of molecular markers
- based on DNA variations, RNA, Epigenome at the single-cell resolution



### Single-Cell from a statistician's perspective



From 10X Genomics

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Let's adopt the ANOVA framework

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•  $Y_{ijr}$  : expression (continuous) for gene *i* in condition *j* at replicate *r* 

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• Perform DE between conditions using model

 $Y_{ijr} \sim \mathcal{N}\left(\mathbb{E}(Y_{ijr}), \sigma^2\right)$ 

 $\mathbb{E}(Y_{ikr}) = \mu_{ij} = \mu + \alpha_i + \beta_j + (\alpha\beta)_{ij}$ 

- The parameters of the model are interpreted as :
  - $\alpha_i$ : mean expression of gene *i* (across conditions),
  - $\beta_j$ : mean expression in condition j (across genes),
  - $(\alpha\beta)_{ii}$ : interaction effect gene x condition
- Allows to integrate normalization while testing

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Testing framework

• Hypothesis : no expression difference between conditions

 $\mathcal{H}_0^i: \{(\alpha\beta)_{i1} = (\alpha\beta)_{i2}\}$ 

• The classical statistic for gene *i* is the Student statistic

$$T_{i} = \frac{|\widehat{\alpha\beta}_{i1} - \widehat{\alpha\beta}_{i2}|}{\widehat{\sigma}} \times \sqrt{2R - 2} \underset{\mathcal{H}_{0}}{\sim} \mathcal{T}(2R - 2)$$

- Estimation of mean fixed effects is done by Maximum Likelihood
- Multiple testing issues are assessed using the FDR control

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What about the estimation of the dispersion parameter ?

- Refinements / difficulties concern the estimation of  $\sigma$ , the dispersion parameter
- A common variance to all genes  $\sigma^2$  : robust but lacks of power
- A specific variance to every gene σ<sub>i</sub><sup>2</sup> : powerful but sensitive to outliers,
  - Large sampling variance
  - To be stabilized empirically
- Groups of variances (combination of both)

### The Generalized Linear Model framework

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•  $Y_{ijr}$ : the read count (positive integer), for gene *i* in condition *j* 

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• Define the Generalized Linear Model (GLM) by setting

 $\begin{array}{lll} Y_{ijr} & \sim & \mathcal{P}(\mu_{ij}) \\ \log \mathbb{E}(Y_{ijr}) = \log(\mu_{ij}) & = & \mu + \alpha_i + \beta_j + (\alpha\beta)_{ij} \end{array}$ 

- Parameters have the same interpretation
- Testing hypotheses are similar :  $\mathcal{H}_0^i$  :  $\{(\alpha\beta)_{i1} = (\alpha\beta)_{i2}\}$
- Dispersion parameter ? Test statistics ?

References

The Exponential family of distributions

Family of distributions that share common mathematical properties

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- The Exponential Family is one of the most widely used
- Consider two types of parameters :

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- $\theta$  the **natural parameter**, related to the location parameter
- $\phi$  the dispersion parameter
- If Y belongs to the exponential family, its density is of the form

$$p(y; \theta, \phi) \propto \exp\left(\frac{y\theta - b(\theta)}{a(\phi)}\right) \qquad \begin{array}{ccc} \theta & \phi \\ \hline \text{Gaussian} & \mu & \sigma^2 \\ \text{Poisson} & \log(\mu) & 1 \\ & \text{Binomial} & \log(\mu) & 1 \end{array}$$

The Generalized Linear Model (GLM)

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• Suppose that  $\mathbb{E}(Y) = \mu$  linearly depends on some covariates X

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$$g(\mathbb{E}(Y)) = g(\mu) = X\beta$$

- $\eta = g(\mu)$  is often called the **linear predictor**
- Most of the time the canonical link is used  $g(\mu) = g(b'(\theta)) = \theta$

	$g(\mu)$	$V(\mu)$
Gaussian	$\mu$	1
Poisson	$\log(\mu)$	$\mu$
Binomial	$logit(\mu)$	$\mu(1-\mu)$

• In GLMs, overdispersion  $\phi$  is not used (exponential dispersion family)

Location / Dispersion relations

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• The first two moments of the exponential family are :

$$\mathbb{E}[Y] = b'( heta) = \mu$$
  
 $\mathbb{V}[Y] = b''( heta) imes a(\phi)$ 

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- The expectation  $\mu$  is a function of  $\theta$  only (location)
- The variance is a function of **both**  $(\theta, \phi)$  (location and dispersion)
- b"(●) is called the variance function (also denoted by V(μ)), and describes how the variance relates to the mean
- Gaussian : (exception!)  $a(\phi) = \sigma^2$  and  $b''(\theta) = 1$  (cst curvature)
- Poisson, Binomial :  $a(\phi) = 1$  (no freedom)

References

### A matter of vocabulary

- Recall that  $\mathbb{V}[Y] = a(\phi) \times V(\mu)$  where  $\phi$  is the dispersion parameter
- The Poisson distribution has no dispersion parameter
- The only possible Discrete Exponential Dispersion model with a disperson parameter are additive models such as **Negative Binomial** or **Poisson-Tweedie**
- Parameter  $\alpha$  may be called dispersion parameter

	$a(\phi)$	$V(\mu)$
Poisson	1	$\mu$
QuasiPoisson	$\phi$	$\mu$
Negative Binomial	1	$\mu + \kappa \mu^2$
Tweedie Poisson	1	$\mu^{{m  ho}}$

Dealing with dispersion estimates

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- If we choose the NB model,  $V(\mu) = \mu + \kappa \mu^2$
- Then we can follow the same procedure compared with the Gaussian case:

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- Use genes as replicates to uncover the mean/variance relationship (one gene=one point)
- perform a regression V(μ) = μ + κ<sub>Global</sub>μ<sup>2</sup> to estimate κ<sub>Global</sub> that would be common to every gene
- In Anders et al., the final "dispersion" estimate is for each gene : max(\(\hat{\alpha}\_{Global}, \(\hat{\alpha}\_i\)) (little loss in power)

Testing Strategies based on LRT

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• Compare different models, for instance

$$log(\mu_{ij}) = \mu + \alpha_i + \beta_j$$
  
$$log(\mu_{ij}) = \mu + \alpha_i + \beta_j + (\alpha\beta)_{ij}$$

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• Use the Ratio of log likelihoods as a Statistics, which incorporates all infos:

$$LRT = -2\log\left(rac{\mathcal{L}(\widehat{\mu},\widehat{lpha},\widehat{eta},\widehat{lpha}ig)}{\mathcal{L}(\widehat{\mu},\widehat{lpha},\widehat{eta})}
ight) \underset{\mathcal{H}_{0}}{\sim} \chi^{2}(\Delta df)$$

This has been shown to be the best strategy on Sequencing data

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### Conclusion: don't think Normal !

- Use **Generalized Linear Models** to perform Count regression, and not Gaussian regression on the log-counts
- Incorporate effects in the model to perform a global analysis that accounts for distributional characteristics
- Do not perform tests that imply Poisson distribution when data are over-dispersed
- Use Likelihood Ratio Tests to compare models
- Overdispersion leads to estimation issues due to **numerical problems**



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How bad is the situation in single cell data ?



Overdispersion is mainly biological because diversity is high between cells [5]

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### Expression is a stochastic bursty process: biological zeros



### The curse of Dropouts

- Low starting amount of RNAs: transcripts will be missed during RT
- Amplification is needed ( $\times 10^6),$  which creates distortions
- Stochasticity of gene expression (bursty process) sparsity of the data, high proportion of zeros
- Dropout depends on cells (different in different wells),
- Lowly expressed genes : sampling / amplification issues
- Highly expressed genes: is more likely to indicate a burst

### Consequences for the analysis of multiple single cells

- The data consist of a snapshot : all cells are not synchronized
- No technical replicate per cell (invasive experiment)
- A lot of zeros in the data : zero inflated count distributions
- For cell r, gene i, condition j, the expression value is modelled by

$$Y_{ijr} \sim \pi_i \delta_0 + (1 - \pi_i) \mathsf{NB}(\mu_{ijr})$$

• Difficulty to discriminate between low expression / no expression

$$\pi_i = f(\mathbb{E}(Y_{ijr}))$$

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### Two-component Generalized Linear Model [11]



### Inference for ZINB models

- Optimization can be based on IRWLS (iterative) combined with the EM algorithm
- EM is used to retrieve the ZI compartment
- Then IRWLS is used to estimate the parameters in the NB compartement
- Quite challenging from the numerical point of view
- Use Bayesian strategies thanks to the Poisson-Gamma representation of NB distributions

$$\mu \sim \Gamma(a, b), \ Y | \mu \sim \mathcal{P}(\mu), \ Y \sim NB(a, \frac{b}{a+1})$$

• Model the sampling process of genes

Normalizing single cell expression data

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• Adjusts for effects related to distributional differences in read counts between cells (sequencing depth)

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• Scaling factor for cell *r* to rescale all cell specific measures on a common scale

$$\mathbb{E}(Y_{ijr}) = s_r \times \mu_{ij}$$

- How to estimate the scaling factor ? RPKM ?
- Library size normalization can be dominated by a handful of highly expressed genes, which can bias downstream analysis .
- Quantile matching ? but difficult to apply with zero Inflation
- Litterature suggests to use the trimmed means proposed by DESeq [13]

### Normalizing single cell expression data [13]

	Cell-specific effects	Gene-specific effects	Not removed by UMIs
Sequencing depth	$\checkmark$		$\checkmark$
Amplification	$\checkmark$	<ul> <li>Image: A set of the set of the</li></ul>	
Capture and RT efficiency	$\checkmark$	<ul> <li>Image: A second s</li></ul>	<b>√</b>
Gene length		<ul> <li>Image: A set of the set of the</li></ul>	
GC content	$\checkmark$	<b>√</b>	✓
mRNA content	$\checkmark$		1

Hypothesis testing becomes more intricate

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- A Compartment model defined by  $(\pi,\mu)$
- Testing hypotheses are similar : H<sup>i</sup><sub>0</sub> : {(π<sub>j</sub>, μ<sub>j</sub>) = (π<sub>j'</sub>, μ<sub>j'</sub>)}
- What is differential expression ? Differential dropout ?
- Difficulty to define  $\mathcal{H}_1$

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### An unprecedented challenge

• Genomics was precursor for data representation and visualization

Publication	cells	tissue	Seq. protocol	clusters
Cadwell et al. (2016)	46	visual cortex	Smart-seq2	2
Tasic et al. (2016)	1,679	visual cortex	SMARTer	49
Macosko et al. (2015)	44,808	retina	Drop-seq	39
10x Genomics	1,306,127	brain cells	10x Gen.Chrom.	39

- We are far beyond the few clusters / some points
- Dimension reduction is mandatory for any analysis (clustering, visualization, GRN inference, etc)

High-dimensional count data

$$x_{ij} = expression of gene j in cell i$$



- **High dimension:** *n* grows but  $\ll p$
- Count data with dropouts

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### A quickie on PCA 1

- PCA represents multivariate data X<sub>n×p</sub> using projection in a lower dimensional space of dimension K < p.</li>
- PCA is non supervised in the sense that the projection is done without guidance
- The strategy is global : there may exist correlations between variables that could be used to summarized the data, and to visualize **X** despite its dimensionality.
- PCA provides orthogonal principal components that best explain the variability of the data globally.
- The key ingredient of PCA is then the empirical covariance matrix

$$\mathbf{S}_{p\times p} = \frac{1}{n-1} \mathbf{X}_c^T \mathbf{X}_c$$

# A guickie on PCA 2

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- PCs  $\mathbf{t}_1, ..., \mathbf{t}_K$ , are defined s.t.  $t_{ik} = \sum_{j=1}^p w_{jk} X_{ij}$
- $w_{jk}$  quantifies the weight of variable  $X_{.j}$  in the constitution of PC k.

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 w<sub>1</sub>, ..., w<sub>K</sub> are determined iteratively by finding the PCs that carry most inertia

$$\mathbb{V}(\mathbf{t}_k) = \mathbf{w}_k^T \mathbf{X}^T \mathbf{X} \mathbf{w}_k.$$

• Solve iteratively the following optimization problem:

$$\mathbf{w}_{k} = \arg \max_{\mathbf{w} \in \mathbb{R}^{p}, \|\mathbf{w}\|=1} \left\{ \mathbf{w}^{T} \mathbf{X}_{c}^{T} \mathbf{X}_{c} \mathbf{w} \right\}, \text{ with } \mathbf{t}_{k} = \mathbf{X} \mathbf{w}_{k} \perp \mathbf{t}_{1}, ... \mathbf{t}_{k-1}.$$

w<sub>1</sub>,...w<sub>r</sub> are the associated eigen vectors of w<sup>T</sup>X<sup>T</sup><sub>c</sub>X<sub>c</sub>w associated with eigen values λ<sub>1</sub>,...λ<sub>r</sub>.

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### A quickie on PCA 3

- A singular value decomposition is the decomposition of a matrix X<sub>c</sub> such that X<sub>c</sub> = UDV<sup>T</sup>,
- $\mathbf{D}_{r \times r} = \text{diag}(\delta_1, ... \delta_r)$  is the diagonal matrix of singular values of  $\mathbf{X}_c$ .
- **U** is orthonormal, whose columns are eigen vectors of  $(\mathbf{X}_{c}\mathbf{X}_{c}^{T})$
- V is orthonormal whose columns are eigen vectors of  $(\mathbf{X}_c^T \mathbf{X}_c)$
- PCA can be rephrased as a minimization problem:

$$\min_{\mathbf{u}\in\mathbb{R}^n,\mathbf{v}\in\mathbb{R}^p,\|\mathbf{u}\|=\|\mathbf{v}\|=1}\|\mathbf{X}_c-\mathbf{U}\mathbf{V}^T\|_F^2$$

• with 
$$\|\mathbf{A}\|_F^2 = \sum_{ij} a_{ij}^2$$
,

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# Matrix factorization: $\mathbf{X} \approx \mathbf{U}\mathbf{V}^T$

 $\begin{array}{ll} \text{Cells:} & \textbf{U} \in \mathbb{R}^{n \times K} \\ \text{Genes:} & \textbf{V} \in \mathbb{R}^{p \times K} \end{array} \right\} \text{ Low dimensional representation}$ 



 $\rightarrow$  Low-rank representation of **X** 

### Matrix factorization: $\mathbf{X} \approx \mathbf{U} \mathbf{V}^{T}$



### **Data visualization:** scatter plot $(u_{i1}, u_{i2})_{i=1:n}$



### Approximation $\mathbf{X} \approx \mathbf{U}\mathbf{V}^T$ ?





### Approximation $\mathbf{X} \approx \mathbf{U}\mathbf{V}^T$ ?



#### **Principal Component Analysis:**

- Find a linear projection of X with maximum variance
- SVD algorithm:  $\underset{\mathbf{U}\in\mathbb{R}^{n\times K},\mathbf{V}\in\mathbb{R}^{p\times K}}{\operatorname{argmin}} \|\mathbf{X}-\mathbf{U}\mathbf{V}^{T}\|_{F}^{2}$
- Least squares approximation

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# RNA-seq data = Counts

Relation between geometry and underlying model  $\|\cdot\|_2 \leftrightarrow \text{Gaussian distribution}$ 

- First idea:  $X_{ij} \sim \mathcal{P}(\lambda)$
- Highly expressed genes
  - $\hookrightarrow \mathsf{large}\ \lambda$
  - $\hookrightarrow \mathsf{Gaussian} \ \mathsf{approximation}$



Figure:  $\mathcal{P}(200)$  empirical distribution

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### RNA-seq data = Counts

Relation between geometry and underlying model  $\|\cdot\|_2 \leftrightarrow$  Gaussian distribution

- First idea:  $X_{ij} \sim \mathcal{P}(\lambda)$
- Highly expressed genes
  - $\hookrightarrow$  large  $\lambda$
  - $\hookrightarrow$  Gaussian approximation



Figure:  $\mathcal{P}(2)$  empirical distribution

# Need for a probabilistic PCA

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- Over-dispersion in RNA-seq data  $\rightarrow Var(X_{ij}) > \mathbb{E}[X_{ij}]$
- Single-cell data: **zero-inflation**  $\rightarrow \mathbb{P}(X_{ij} = 0) > e^{-\lambda}$

#### Embed PCA with a probabilistic model

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- $X_{ij} \sim$  probability distribution in the exponential family
- Factorization of  $\mathbb{E}[\textbf{X}]$  rather than X
- Replace  $\|\cdot\|_2$  approximation by likelihood-based approaches

References

Generalized PCA[3] and Poisson NMF [8]

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- $X_{ij} \sim \mathcal{P}(\lambda_{ij})$  with the Poisson rate matrix  $\mathbf{\Lambda} = [\lambda_{ij}]_{n \times p}$
- Decompose  $\mathbb{E}[\mathbf{X}] = \mathbf{\Lambda}$  such that  $\lambda_{ij} = \sum_k U_{ik} V_{kj}$



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Probabilistic PCA with a ZI-Gamma-Poisson model [1]

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- Factors U,V become Gamma latent variables
- Marginal distribution is over-dispersed: Var(X<sub>ij</sub>) > ℝ[X<sub>ij</sub>]
- $D_{ij} = \text{drop-out event indicator}$

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Probabilistic variable selection with a spike and slab model

• Sparsity-inducing priors:

$$V_{jk} \sim \pi_S \delta_0 + (1 - \pi_S) \Gamma(\beta_{k,1}, \beta_{k,2})$$



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### Model inference

- Recover the posterior distributions  $\boldsymbol{U} \,|\, \boldsymbol{X}$  and  $\boldsymbol{V} \,|\, \boldsymbol{X}$
- Estimate the factors as  $\widehat{\mathbf{U}} = \mathbb{E}[\mathbf{U} \,|\, \mathbf{X}]$  and  $\widehat{\mathbf{V}} = \mathbb{E}[\mathbf{V} \,|\, \mathbf{X}]$
- Posteriors are not explicit
- Variational inference: approximation of the posteriors

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### Beyond Linear projections

- Linear methods are powerful for planar structures
- High dimensional datasets are characterized by multiscale properties (local / global structures)
- May not be the most powerful for manifolds
- Non Linear projection methods aim at preserving local characteristics of distances



Stochastic Neighbor Embedding 1 [14]

•  $(x_1, \ldots, x_n)$  are the points in the high dimensional space  $\mathbb{R}^p$ ,

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• Consider a similarity between points:

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$$p_{i|j} = \frac{\exp(-\|x_i - x_j\|^2 / 2\sigma_i^2)}{\sum_{k \neq i} \exp(-\|x_k - x_j\|^2 / 2\sigma_k^2)}, \ p_{ij} = (p_{i|j} + p_{j|i}) / 2N$$

- $\sigma$  smooths the data (linked to the regularity of the target manifold)
- σ is chosen such that the entropy of p is fixed to a given value of the so-called perplexity

$$\exp\left(-\sum_{ij}p_{ij}\log(p_{ij})\right)$$

### Visual inspection of the influence of $\sigma$ [7]



### Stochastic Neighbor Embedding 2

- Consider  $(y_1, \ldots, y_n)$  are points in the low dimensional space  $\mathbb{R}^2$
- Consider a similarity between points in the new representation:

$$q_{i|j} = \frac{\exp(-\|y_i - y_j\|^2)}{\sum_{k \neq i} \exp(-\|y_k - y_j\|^2)}$$

• Robustify this kernel by using Student(1) kernels (ie Cauchy)

$$q_{i|j} = \frac{(1 + \|y_i - y_j\|^2)^{-1}}{\sum_{k \neq i} (1 + \|y_i - y_k\|^2)^{-1}}$$

### Stochastic Neighbor Embedding 3

• Minimize the KL between *p* and *q* so that the data representation minimizes:

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$$\mathcal{C}(y) = \sum_{ij} \mathcal{KL}(p_{ij}, q_{ij})$$

The cost function is not convex

$$\frac{\partial C(y)}{\partial y} = \sum_{ij} (p_{ij} - q_{ij})(y_i - y_j)$$

• Very sensitive to starting values

### Stochastic Neighbor Embedding 4 [7]



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## Stochastic Neighbor Embedding 5 [7]



### Properties of t-SNE

- good at preserving local distances (intra-cluster variance)
- not so good for global representation (inter-cluster variance)
- hence good at creating clusters of points that are close, but bad at positionning clusters wrt each other
- preprocessing very important : initialize with PCA and feature selection plus log transform (non linear transform)
- percent of explained variance ? interpretation of the q distribution ?

### Single-cell RNAseq example[9]



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### A taxonomy of Dimension Reduction Methods [4]



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# Conclusions of a comparative study [4]

- local methods suffer from the choice of the smoothing (neighborhood) parameter
- Kernel PCA suffers from the choice of the Kernel to correctly approximate the manifold.
- Setting the optimization problem is the key (convex or not), trivial solutions, local optima, computationally feasible
- nonlinear techniques for dimensionality reduction are, despite their large variance, often not capable of outperforming traditional linear techniques such as PCA.

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### What is missing ?

- Clustering
- Trajectory inference
- Epigenomics
- Networks

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