Inferring multiple graph structures

Julien Chiquet, jointly with Christophe Ambroise, Camille Charbonnier, Yves Grandvalet, Catherine Matias...

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INRA Toulouse - 21 Janvier 2011



Chiquet, Grandvalet, Ambroise, Statistics and Computing, 2010.

Inferring multiple graphical structures.

Chiquet, Grasseau, Charbonnier and Ambroise, New release of R-package SIMoNe.

http://stat.genopole.cnrs.fr/softwares/simone





few arrays ⇔ few examples lots of genes ⇔ high dimension interactions ⇔ very high dimension

Which interactions?

The main trouble is the low sample size and high dimensional setting

Our main hope is to benefit from sparsity: few genes interact

Merge several experimental conditions experiment 1 experiment 2





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Inferring each graph independently does not help experiment 1 experiment 2 exp







By pooling all the available data experiment 1 experiment 2



experiment 1



experiment 2







By breaking the separability experiment 1 experiment 2



By breaking the separability experiment 1 experiment 2



Statistical model

- Multi-task learning
- Geometrical insights
- Optimization strategy
- Theoretical results

Experiments

Outline

Statistical model

- Multi-task learning
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- Optimization strategy
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- Experiments

Let

- $X = (X_1, \dots, X_p) \sim \mathcal{N}(\mathbf{0}_p, \mathbf{\Sigma})$ and assume n i.i.d. copies of X,
- X be the $n \times p$ matrix whose kth row is X_k ,
- $\boldsymbol{\Theta} = (\theta_{ij})_{i,j \in \mathcal{P}} \triangleq \boldsymbol{\Sigma}^{-1}$ be the concentration matrix.

$\begin{aligned} & \text{Graphical interpretation} \\ & \text{Since } \operatorname{cor}_{ij|\mathcal{P}\setminus\{i,j\}} = -\theta_{ij}/\sqrt{\theta_{ii}\theta_{jj}} \text{ for } i \neq j, \\ & X_i \perp X_j | X_{\mathcal{P}\setminus\{i,j\}} \Leftrightarrow \begin{cases} & \theta_{ij} = 0 \\ & \text{or} \\ & \text{edge } (i,j) \notin \text{ network.} \end{cases} \end{aligned}$

 \rightsquigarrow non zeroes in Θ describes the graph structure.

Let $\mathbf{S} = n^{-1} \mathbf{X}^{\mathsf{T}} \mathbf{X}$ be the empirical variance-covariance matrix: \mathbf{S} is a sufficient statistic for $\mathbf{X} \Rightarrow \mathcal{L}(\mathbf{\Theta}; \mathbf{X}) = \mathcal{L}(\mathbf{\Theta}; \mathbf{S})$

The log-likelihood

$$\mathcal{L}(\boldsymbol{\Theta}; \mathbf{S}) = \frac{n}{2} \log \det(\boldsymbol{\Theta}) - \frac{n}{2} \operatorname{trace}(\mathbf{S}\boldsymbol{\Theta}) - \frac{n}{2} \log(2\pi).$$

The MLE of Θ is \mathbf{S}^{-1}

 \frown not defined for n < p

 \frown not sparse \Rightarrow fully connected graph

Inferring multiple graph structures

Penalized Approaches

Penalized Likelihood (Banerjee et al., 2008)

$$\underset{\boldsymbol{\Theta} \in \mathbb{S}_{+}}{\operatorname{maximize}} \mathcal{L}(\boldsymbol{\Theta}; \mathbf{S}) - \lambda \|\boldsymbol{\Theta}\|_{1}$$

 \smile well defined for n < p

- \sim sparse \Rightarrow sensible graph
- \frown SDP of size $\mathcal{O}(p^2)$ (solved by Friedman *et al.*, 2007)

Neighborhood Selection (Meinshausen & Bülhman, 2006) $\widehat{\beta} = \underset{\beta \in \mathbb{R}^{p-1}}{\operatorname{argmin}} \frac{1}{n} \| \mathbf{X}_j - \mathbf{X}_{\setminus j} \beta \|_2^2 + \lambda \| \beta \|_1$ where \mathbf{X}_j is the *j*th column of \mathbf{X} and $\mathbf{X}_{\setminus j}$ is \mathbf{X} deprived of \mathbf{X}_j

 $\sim p$ independent LASSO problems of size (p-1)

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where \mathbf{X}_j is the *j*th column of \mathbf{X} and $\mathbf{X}_{\setminus j}$ is \mathbf{X} deprived of \mathbf{X}_j

- not symmetric, not positive-definite
- -p independent LASSO problems of size (p-1)

Neighborhood vs. Likelihood

Pseudo-likelihood (Besag, 1975) $\mathbb{P}(X_1, \dots, X_p) \simeq \prod_{j=1}^p \mathbb{P}(X_j | \{X_k\}_{k \neq j})$ $\widetilde{\mathcal{L}}(\Theta; \mathbf{S}) = \frac{n}{2} \log \det(\mathbf{D}) - \frac{n}{2} \operatorname{trace} \left(\mathbf{S}\mathbf{D}^{-1}\mathbf{\Theta}^2\right) - \frac{n}{2} \log(2\pi)$ $\mathcal{L}(\Theta; \mathbf{S}) = \frac{n}{2} \log \det(\Theta) - \frac{n}{2} \operatorname{trace}(\mathbf{S}\Theta) - \frac{n}{2} \log(2\pi)$ with $\mathbf{D} = \operatorname{diag}(\mathbf{\Theta})$.

Proposition (Ambroise, Chiquet, Matias, 2008) Neighborhood selection leads to the graph maximizing the penalized pseudo-log-likelihood

Proof: $\hat{\beta}_i = -\frac{\theta_{ij}}{\overline{\theta}_{jj}}$, where $\hat{\Theta} = \arg \max_{\Theta} \hat{\mathcal{L}}(\Theta; \mathbf{S}) - \lambda \|\Theta\|_1$

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We have T samples (experimental cond.) of the same variables

- $\mathbf{X}^{(t)}$ is the t^{th} data matrix, $\mathbf{S}^{(t)}$ is the empirical covariance
- examples are assumed to be drawn from $\mathcal{N}(\mathbf{0}, \mathbf{\Sigma}^{(t)})$

Ignoring the relationships between the tasks leads to separable objectives

$$\underset{\mathbf{\Theta}^{(t)} \in \mathbb{R}^{p \times p}, t=1...,T}{\text{maximize}} \widetilde{\mathcal{L}}(\mathbf{\Theta}^{(t)}; \mathbf{S}^{(t)}) - \lambda \|\mathbf{\Theta}^{(t)}\|_{1}$$

Multi-task learning = solving the T tasks jointly We may **couple** the objectives

- through the fitting term term,
- through the penalty term.

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Coupling through the fitting term

Intertwined LASSO

$$\underset{\boldsymbol{\Theta}^{(t)},t...,T}{\text{maximize}} \sum_{t=1}^{T} \widetilde{\mathcal{L}}(\boldsymbol{\Theta}^{(t)}; \widetilde{\mathbf{S}}^{(t)}) - \lambda \| \boldsymbol{\Theta}^{(t)} \|_{1}$$

S̄ = 1/n ∑_{t=1}^T n_tS^(t) is the "pooled-tasks" covariance matrix.
 S̃^(t) = αS^(t) + (1 − α)S̄ is a mixture between specific and pooled covariance matrices.

- $\alpha = 0$ pools the data sets and infers a single graph
- $\alpha = 1$ separates the data sets and infers T graphs independently
- $\alpha = 1/2$ in all our experiments

We group parameters by sets of corresponding edges across graphs:



Graphical group-LASSO

$$\underset{\boldsymbol{\Theta}^{(t)},t...,T}{\text{maximize}} \sum_{t=1}^{T} \widetilde{\mathcal{L}}\left(\boldsymbol{\Theta}^{(t)}; \mathbf{S}^{(t)}\right) - \lambda \sum_{i \neq j} \left(\sum_{t=1}^{T} \left(\theta_{ij}^{(t)}\right)^{2}\right)^{1/2}$$

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- Same grouping, and bet that correlations are likely to be sign consistent
- Gene interactions are either inhibitory or activating across assays

Graphical cooperative-LASSO



$$\underset{\substack{\boldsymbol{\Theta}^{(t)}\\t=1,\dots,T}}{\operatorname{maximize}} \sum_{t=1}^{T} \widetilde{\mathcal{L}}(\mathbf{S}^{(t)}; \boldsymbol{\Theta}^{(t)}) - \lambda \sum_{i \neq j} \left\{ \left(\sum_{t=1}^{T} \left[\theta_{ij}^{(t)} \right]_{+}^{2} \right)^{\frac{1}{2}} + \left(\sum_{t=1}^{T} \left[\theta_{ij}^{(t)} \right]_{-}^{2} \right)^{\frac{1}{2}} \right\}$$

where $[u]_+ = \max(0, u)$ and $[u]_- = \min(0, u)$.

Plausible in many other situations

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Experiments

A Geometric View of Sparsity

Constrained Optimization



$$\max_{\beta_1,\beta_2} \mathcal{L}(\beta_1,\beta_2) - \lambda \Omega(\beta_1,\beta_2)$$

A Geometric View of Sparsity

Constrained Optimization



$$\begin{array}{l} \max_{\beta_1,\beta_2} \mathcal{L}(\beta_1,\beta_2) - \lambda \Omega(\beta_1,\beta_2) \\ \Leftrightarrow & \left\{ \begin{array}{l} \max_{\beta_1,\beta_2} & \mathcal{L}(\beta_1,\beta_2) \\ \text{s.t.} & \Omega(\beta_1,\beta_2) \leq c \end{array} \right. \end{array}$$

 $\mathcal{L}(eta_1,eta_2)$

A Geometric View of Sparsity

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- the set is contained in one half-space
- the set has at least one point on the hyperplane



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There are Supporting Hyperplane at all points of convex sets: Generalize tangents

A Geometric View of Sparsity

Generalizes normals



A Geometric View of Sparsity

Generalizes normals



A Geometric View of Sparsity

Generalizes normals



A Geometric View of Sparsity Dual Cone

Generalizes normals



Shape of dual cones \Rightarrow sparsity pattern

Admissible set

- ▶ 2 tasks (T = 2)
- 2 coefficients (p = 2)

$$\sum_{i=1}^{2} \left(\sum_{t=1}^{2} \beta_{i}^{(t)^{2}}\right)^{1/2} \leq 1$$



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Estimate the j^{th} neighborhood of the T graphs

$$\underset{\boldsymbol{\Theta}^{(t)},t=1\dots,T}{\text{maximize}} \sum_{t=1}^{T} \tilde{\mathcal{L}}(\boldsymbol{\Theta}^{(t)}; \mathbf{S}^{(t)}) - \lambda \ \Omega(\mathbf{K}^{(t)})$$

decomposes into p convex optimization problems of size

$$\widehat{\boldsymbol{\beta}}_{j} = \operatorname*{argmin}_{\boldsymbol{\beta} \in \mathbb{R}^{T \times (p-1)}} f_{j}(\boldsymbol{\beta}) + \lambda \; \Omega(\boldsymbol{\beta})$$

where $\widehat{\beta}_j$ is a minimizer iff $0 \in \nabla_{\beta} f_j(\beta) + \lambda \partial_{\beta} \Omega(\beta)$

Estimate the j^{th} neighborhood of the T graphs

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$$\Omega(\boldsymbol{\beta}) = \sum_{t=1}^{T} \left\| \boldsymbol{\beta}^{(t)} \right\|_{1} ,$$

where
$$\boldsymbol{\beta} = \left(\boldsymbol{\beta}^{(1)}, \dots, \boldsymbol{\beta}^{(T)}\right)^{\mathsf{T}}$$
, $\boldsymbol{\beta}^{(t)} \in \mathbb{R}^{p-1}$

Inferring multiple graph structures

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$$\Omega(\boldsymbol{\beta}) = \sum_{i=1}^{p-1} \left\| \boldsymbol{\beta}_i^{[1:T]} \right\|_2$$

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$$\Omega(\boldsymbol{\beta}) = \sum_{i=1}^{p-1} \left(\left\| \left(\boldsymbol{\beta}_i^{[1:T]} \right)_+ \right\|_2 + \left\| \left(- \boldsymbol{\beta}_i^{[1:T]} \right)_+ \right\|_2 \right)$$

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Inferring multiple graph structures

Active set algorithm: • yellow belt



Active set algorithm: orange belt

```
// 0. INITIALIZATION \beta \leftarrow 0, \mathcal{A} \leftarrow \emptyset
while \mathbf{0} \notin \partial_{\boldsymbol{\beta}} L(\boldsymbol{\beta}) do
      // 1. MASTER PROBLEM: OPTIMIZATION WITH RESPECT TO \beta_A
      Find a solution \mathbf{h} to the smooth problem
       \boldsymbol{\beta}_{A} \leftarrow \boldsymbol{\beta}_{A} + \mathbf{h}
      // 2. IDENTIFY NEWLY ZEROED VARIABLES;
      \mathcal{A} \leftarrow \mathcal{A} \setminus \{i\}
      // 3. IDENTIFY NEW NON-ZERO VARIABLES:
      // Select a candidate i \in \mathcal{A}^c which violates the more the optimality
      conditions
      if it exists such an i then
             \mathcal{A} \leftarrow \mathcal{A} \cup \{i\}
      else
             Stop and return \beta, which is optimal
      end
end
```

Active set algorithm: green belt

// 0. INITIALIZATION $\beta \leftarrow 0, \mathcal{A} \leftarrow \emptyset$ while $\mathbf{0} \notin \partial_{\boldsymbol{\beta}} L(\boldsymbol{\beta})$ do // 1. MASTER PROBLEM: OPTIMIZATION WITH RESPECT TO β_A Find a solution \mathbf{h} to the smooth problem $\nabla_{\mathbf{h}} f(\boldsymbol{\beta}_{A} + \mathbf{h}) + \lambda \partial_{\mathbf{h}} \Omega(\boldsymbol{\beta}_{A} + \mathbf{h}) = 0, \text{ where } \partial_{\mathbf{h}} \Omega = \{\nabla_{\mathbf{h}} \Omega\}.$ $\boldsymbol{\beta}_{A} \leftarrow \boldsymbol{\beta}_{A} + \mathbf{h}$ // 2. IDENTIFY NEWLY ZEROED VARIABLES; while $\exists i \in \mathcal{A} : \beta_i = 0$ and $\min_{\nu \in \partial_{\beta_i} \Omega} \left| \frac{\partial f(\beta)}{\partial \beta_i} + \lambda \nu \right| = 0$ do $| \quad \mathcal{A} \leftarrow \mathcal{A} \setminus \{i\}$ end // 3. IDENTIFY NEW NON-ZERO VARIABLES; // Select a candidate $i \in \mathcal{A}^c$ such that an infinitesimal change of β_i provides the highest reduction of L $i \leftarrow rgmax_{j \in \mathcal{A}^c} v_j$, where $v_j = \min_{\nu \in \partial_{\beta_i} \Omega} \left| rac{\partial f(oldsymbol{eta})}{\partial eta_j} + \lambda
u
ight|$ if $v_i \neq 0$ then $\mathcal{A} \leftarrow \mathcal{A} \cup \{i\}$ else Stop and return β , which is optimal end end

Inferring multiple graph structures

Statistical model

- Multi-task learning
- Geometrical insights
- Optimization strategy
- Theoretical results

Experiments

(Sparse) linear regression setup

Let Y be a response variable, $X = (X_1, ..., X_p)$ a vector of p features,

$$Y = X\beta^{\star} + \boldsymbol{\varepsilon} = \sum_{j=1}^{p} X_{j}\beta_{j}^{\star} + \boldsymbol{\varepsilon}, \quad \boldsymbol{\varepsilon} \sim \mathcal{N}(\mathbf{0}, \sigma^{2}\mathbf{I}) ,$$

• $S = \{j, \beta_j^* \neq 0\}$ is the true support,

• β^* has a group structure $\{\mathcal{G}_k\}_{k=1,\dots,K}$.

Cooperative-Lasso estimate of β^*

Given the training vector $\mathbf{y} = (y_1, \dots, y_n)^{\mathsf{T}}$ and the $n \times p$ design matrix \mathbf{X} whose *j*th column $\mathbf{x}_j = (x_j^1, \dots, x_j^n)^{\mathsf{T}}$,

$$\hat{\boldsymbol{\beta}}^{\text{coop}} = \operatorname*{arg\,min}_{\boldsymbol{\beta} \in \mathbb{R}^p} \frac{1}{2} \left\| \mathbf{y} - \mathbf{X} \boldsymbol{\beta} \right\|_n^2 + \lambda_n \sum_{k=1}^K \left\| [\boldsymbol{\beta}_{\mathcal{G}_k}]_+ \right\| + \left\| [\boldsymbol{\beta}_{\mathcal{G}_k}]_- \right\|,$$

Inferring multiple graph structures

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Inferring multiple graph structures

Let $\Psi = \mathbb{E}XX^{\intercal}$ be the covariance matrix of X.

- (A1) X and Y have finite fourth order moments $\mathbb{E}||X||^4 < \infty$, $\mathbb{E}||Y||^4 < \infty$,
- (A2) the covariance matrix $\Psi = \mathbb{E}XX^{\intercal} \in \mathbb{R}^{p \times p}$ is invertible,
- (A3) for every k = 1, ..., K, if $\|[\beta^*]_+\| > 0$ and $\|[\beta^*]_-\| > 0$ then for every $j \in \mathcal{G}_k \ \beta_j^* \neq 0$. (There should not be any zero in a group with positive and negative coefficients).

Irrepresentability condition

Define $\mathcal{S}_k = \mathcal{S} \cap \mathcal{G}_k$ the support within a group and

 $D(\boldsymbol{\beta})]_{jj} = \|[\operatorname{sign}(\beta_j)\boldsymbol{\beta}_{\mathcal{G}_k}]_+\|^{-1}.$

Assume there exists $\eta > 0$ such that

► for every group k to switch off (where $S_k^c = \mathcal{G}_k$), $\max(\|[\Psi_{S_k^c S} \Psi_{SS}^{-1} D(\beta_S^*) \beta_S^*]_+\|, \|[\Psi_{S_k^c S} \Psi_{SS}^{-1} D(\beta_S^*) \beta_S^*]_-\|) \le 1 - \eta,$

► for every group k with zero coefficients and either positive or negative coefficients, define $\nu_k = 1$ if positive coefficients are activated, $\nu_k = -1$ otherwise, and require

$$\begin{cases} \nu_k \Psi_{\mathcal{S}_k^c \mathcal{S}} \Psi_{\mathcal{S} \mathcal{S}}^{-1} D(\boldsymbol{\beta}_{\mathcal{S}}^{\star}) \boldsymbol{\beta}_{\mathcal{S}}^{\star} \leq 0 \quad \text{component-wise} \\ \|\Psi_{\mathcal{S}_k^c \mathcal{S}} \Psi_{\mathcal{S} \mathcal{S}}^{-1} D(\boldsymbol{\beta}_{\mathcal{S}}^{\star}) \boldsymbol{\beta}_{\mathcal{S}}^{\star}\| \leq 1 - \eta. \end{cases}$$

Theorem (Chiquet, Grandvalet, Charbonnier, in progress!)

If assumptions (A1-3) are satisfied and if there exists $\eta > 0$, then for every sequence λ_n such that $\lambda_n = \lambda_0 n^{-\gamma}, \ \gamma \in]0, 1/2[$,

$$\hat{\boldsymbol{\beta}}^{\text{coop}} \xrightarrow{P} \boldsymbol{\beta}^{\star} \text{ and } \mathbb{P}(\mathcal{S}(\hat{\boldsymbol{\beta}}^{\text{coop}}) = \mathcal{S}) \to 1.$$
 (1)

Asymptotically, the cooperative-Lasso is unbiased and enjoys exact support recovery (even when there are irrelevant variables within a group G_k).

Statistical model

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Experiments

Data Generation

We set

- the number of nodes p
- the number of edges K
- the number of examples n

Process

- 1. Generate a random adjacency matrix with 2 K off-diagonal terms
- 2. Compute the normalized Laplacian L
- 3. Generate a symmetric matrix of random signs ${f R}$
- 4. Compute the concentration matrix $\Theta_{ij}^{\star} = L_{ij} R_{ij}$
- 5. compute Σ^* by pseudo-inversion of Θ^*
- 6. generate correlated Gaussian data $\sim \mathcal{N}(\mathbf{0}, \Sigma^{\star})$

Simulating Related Tasks

Generate

- 1. an "ancestor" with p = 20 nodes and K = 20 edges
- 2. T = 4 children by adding and deleting δ edges
- 3. T = 4 Gaussian samples



Figure: ancestor and children with $\delta = 2$ perturbations
Simulating Related Tasks

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Figure: ancestor and children with $\delta = 2$ perturbations

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Figure: ancestor and children with $\delta = 2$ perturbations

Precision/Recall curve precision = TP/(TP+FP) recall = TP/P (power)

Simulation results large sample size



Figure: $n_t = 100, \delta = 1$

Simulation results large sample size



Figure: $n_t = 100, \, \delta = 3$

Simulation results large sample size



Figure: $n_t = 100$, $\delta = 5$

Simulation results medium sample size



Figure: $n_t = 50, \delta = 1$

Simulation results medium sample size



Figure: $n_t = 50$, $\delta = 3$

Simulation results medium sample size



Figure: $n_t = 50$, $\delta = 5$

Simulation results small sample size



Figure: $n_t = 25$, $\delta = 1$

Simulation results small sample size



Figure: $n_t = 25$, $\delta = 3$

Simulation results small sample size



Figure: $n_t = 25$, $\delta = 5$

Two types of patients

Patient response can be classified either as

- 1. pathologic complete response (PCR)
- 2. residual disease (not PCR)

Gene expression data

- 133 patients (99 not PCR, 34 PCR)
- 26 identified genes (differential analysis)

Package Demo

cancer data: Coop-Lasso

Jeanmouin, Guedj, Ambroise (preprint http://arxiv.org) Defining a robust biological prior from Pathway Analysis to drive Network Inference

Marine will speak at SMPGD '11 🙂

"Due to the vast space of possible networks and the relatively small amount of data available, inferring genetic networks from gene expression data is one of the most challenging work in the post-genomic era. (...) We propose an original approach for inferring gene regulation network using a robust biological prior on structure in order to limit the set of candidate networks."

- Clarified links between neighborhood selection and graphical LASSO
- Identified the relevance of Multi-Task Learning in network inference
- First methods for inferring multiple Gaussian Graphical Models
- Consistent improvements upon the available baseline solutions
- Available in the R package SIMoNe

Perspectives

Issues

- 1. How can we choose for a unique network? (should we?)
 - Explore model-selection capabilities,
 - Network comparison.
- 2. Robustness
 - Test the validity of an edge ? Of a whole motif ?
 - Bootstrap greatly improves the inference but is computationally intensive,
 - Introduce more biological prior (semi-supervised learning).
- 3. Biological studies
 - Breast cancer (Marine),
 - Parkinson (with J.-C. Corvol, Pitié Salpétrière and Camille),
 - Bacillus subtilis and Staphylococcus aureus (ANR NOUGA déposée: heterogeneous data, RNAseq, new, prior etc.).

Coop-Lasso

Theoretical analysis and other applications in genetics with penalized linear / logistic regression.

Model selection

More details on optimisation

Theory based penalty choices

- 1. Optimal order of penalty in the $p \gg n$ framework: $\sqrt{n \log p}$ Bunea et al. 2007, Bickel et al. 2009
- 2. Control on the probability of connecting two distinct connectivity sets

Meinshausen et al. 2006, Banerjee et al. 2008, Ambroise et al. 2009

→ practically much too conservative

Cross-validation

- Optimal in terms of prediction, not in terms of selection
- Problematic with small samples: changes the sparsity constraint due to sample size

Tuning the penalty parameter

Theorem (Zou et al. 2008)

$$\mathrm{df}(\hat{\beta}_{\lambda}^{\mathsf{lasso}}) = \left\| \hat{\beta}_{\lambda}^{\mathsf{lasso}} \right\|_{0}$$

Straightforward extensions to the graphical framework

BIC(
$$\lambda$$
) = $\mathcal{L}(\hat{\Theta}_{\lambda}; \mathbf{X}) - df(\hat{\Theta}_{\lambda}) \frac{\log n}{2}$

$$\operatorname{AIC}(\lambda) = \mathcal{L}(\hat{\Theta}_{\lambda}; \mathbf{X}) - \operatorname{df}(\hat{\Theta}_{\lambda})$$

Rely on asymptotic approximations, but still relevant for small data set

Model selection

More details on optimisation

Decomposition strategy (1)

Consider the $(pT) \times (pT)$ block-diagonal matrix C composed by the empirical covariance matrices of each tasks

$$\mathbf{C} = \begin{pmatrix} \mathbf{S}^{(1)} & 0 \\ & \ddots & \\ 0 & & \mathbf{S}^{(T)} \end{pmatrix},$$

and define

$$\mathbf{C}_{\backslash i \backslash i} = \begin{pmatrix} \mathbf{S}_{\backslash i \backslash i}^{(1)} & 0 \\ & \ddots & \\ 0 & & \mathbf{S}_{\backslash i \backslash i}^{(T)} \end{pmatrix}, \ \mathbf{C}_{i \backslash i} = \begin{pmatrix} \mathbf{S}_{i \backslash i}^{(1)} \\ \vdots \\ \mathbf{S}_{i \backslash i}^{(T)} \end{pmatrix}$$

The $(p-1)T \times (p-1)T$ matrix $\mathbf{C}_{\setminus i \setminus i}$ is the matrix \mathbf{C} where we removed each line and each column pertaining to variable *i*.

Decomposition strategy (2)

Estimate the i^{th} neighborhood of the T tasks bind together

$$\underset{\boldsymbol{\Theta}^{(t)},t=1\dots,T}{\operatorname{argmax}} \sum_{t=1}^{T} \tilde{\mathcal{L}}(\boldsymbol{\Theta}^{(t)}; \mathbf{S}^{(t)}) - \lambda \ \Omega(\boldsymbol{\Theta}^{(t)})$$

decomposes into p convex optimization problems

$$\operatorname*{argmin}_{\boldsymbol{\beta} \in \mathbb{R}^{T \times (p-1)}} f(\boldsymbol{\beta}; \mathbf{C}) + \lambda \ \Omega(\boldsymbol{\beta}),$$

where we set $oldsymbol{eta}^{(t)} = oldsymbol{\Theta}^{(t)}_{i \setminus i}$ and

$$\boldsymbol{\beta} = \begin{pmatrix} \boldsymbol{\beta}^{(1)} \\ \vdots \\ \boldsymbol{\beta}^{(T)} \end{pmatrix} \in \mathbb{R}^{T \times (p-1)}.$$

$$\min_{\boldsymbol{\beta} \in \mathbb{R}^{T \times (p-1)}} L(\boldsymbol{\beta}) = f(\boldsymbol{\beta}) + \Omega(\boldsymbol{\beta}) \ ,$$

 $oldsymbol{eta}$ is a minimizer lif $\mathbf{0}_p\in\partial_{oldsymbol{eta}}L(oldsymbol{eta})$, with

$$\partial_{\boldsymbol{\beta}} L(\boldsymbol{\beta}) = \nabla_{\boldsymbol{\beta}} f(\boldsymbol{\beta}) + \lambda \partial_{\boldsymbol{\beta}} \Omega(\boldsymbol{\beta}).$$

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For the graphical Intertwined LASSO

$$\Omega(\boldsymbol{\beta}) = \sum_{t=1}^{T} \left\| \boldsymbol{\beta}^{(t)} \right\|_{1} ,$$

where the grouping effect is managed by the function f.

$$\min_{\boldsymbol{\beta} \in \mathbb{R}^{T \times (p-1)}} L(\boldsymbol{\beta}) = f(\boldsymbol{\beta}) + \Omega(\boldsymbol{\beta}) \ ,$$

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$$\partial_{\boldsymbol{\beta}} L(\boldsymbol{\beta}) = \nabla_{\boldsymbol{\beta}} f(\boldsymbol{\beta}) + \lambda \partial_{\boldsymbol{\beta}} \Omega(\boldsymbol{\beta}).$$

For the graphical Group-LASSO

$$\Omega(\boldsymbol{\beta}) = \sum_{i=1}^{p-1} \left\| \boldsymbol{\beta}_i^{[1:T]} \right\|_2 \; ,$$

where $\boldsymbol{\beta}_i^{[1:T]} = \left(\beta_i^{(1)}, \dots, \beta_i^{(T)}\right)^{\mathsf{T}} \in \mathbb{R}^T$ is the vector of the *i*th component across tasks.

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$$\partial_{\boldsymbol{\beta}} L(\boldsymbol{\beta}) = \nabla_{\boldsymbol{\beta}} f(\boldsymbol{\beta}) + \lambda \partial_{\boldsymbol{\beta}} \Omega(\boldsymbol{\beta}).$$

For the graphical Coop-LASSO

$$\Omega(\boldsymbol{\beta}) = \sum_{i=1}^{p-1} \left(\left\| \left[\boldsymbol{\beta}_i^{[1:T]} \right]_+ \right\|_2 + \left\| \left[\boldsymbol{\beta}_i^{[1:T]} \right]_- \right\|_2 \right) \right\|_2$$

where $\boldsymbol{\beta}_i^{[1:T]} = \left(\beta_i^{(1)}, \dots, \beta_i^{(T)}\right)^{\mathsf{T}} \in \mathbb{R}^T$ is the vector of the *i*th component across tasks.